

Polarimetric Target Detection Using Statistic of the Degree of Polarization

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Abstract—The degree of polarization (DoP) can be utilized as a detection statistic in the polarimetric radar to achieve target detection performance improvement. In this paper, a polarimetric radar model is established, which includes reflections from both target and clutter at first. Then, probability density functions (PDFs) of the estimated DoP are expressed in closed form, which is derived from joint eigenvalue distributions of complex noncentral Wishart matrices. The detector is developed and evaluated theoretically on the basis of the statistical properties of the DoP. Finally, a comparison between the new DoP detector and single-polarization detector is presented against real data. The performance improvement is demonstrated by the comparison results.

1. INTRODUCTION

In recent years, most radar systems have the ability to transmit and receive electromagnetic waves at orthogonal polarization (i.e., horizontal and vertical). The polarization information has become an important tool for improving target detection performance from the clutter and jammer background [1, 2].

Since the optimum polarization detector has been designed by Novak et al. in 1989 [2], many polarimetric detection algorithms have been subsequently developed [3–5]. Most of these detectors are based on the Kelly's generalized likelihood ratio (GLR) test [6], which assumes that the clutter covariance can be well estimated by the training data or can be obtained as priori knowledge. This may not be reasonable, especially when the targets exist in nonhomogeneous clutter environment. Moreover, the full polarization scattering matrix (PSM) has to be measured in order to carry on those detectors. However, lots of radar systems can only operate at a dual-polarization mode, at which a single polarization is transmitted, and two orthogonal polarizations (i.e., denoted HH and HV) are received. To overcome the above problems, an important variable available to characterize polarization for polarimetric radar shall be considered as a new detection statistic in this paper, which is the degree of polarization (DoP).

In studies of partially polarized planar waves, the DoP is defined as the ratio of the power of the waves' fully polarized components to its total power and is obtained from Wolf's polarization covariance matrix (PCM) [7, 8]. It has been demonstrated that the physical meaning of the DoP can always be preserved [9, 10]. Furthermore, as an eigenvalue-derived variable, the DoP is independent of the received polarimetric channels chosen to sample the electromagnetic wave.

In practice, the real DoP must be estimated via finite radar received vector samples, which are normally described as random variables because of the random nature of the partially polarized waves. The applications of the DoP statistical properties have already been widely carried out in radar research, geophysics, and optics [11–14]. Specifically, it is found that the mean of the estimated DoP can provide discriminating information for various terrain types such as urban, vegetation, and ocean during the Pol-SAR data research in [11]. Moreover, the estimated DoP can be used as test statistics to detect

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the man-made metallic objects (including high-voltage transmission towers, ships, buoys and so on) in clutter environments or to monitor oil spills on the surface of the ocean [12]. In this paper, we mainly focus on the performance improvement of the DoP detector with respect to the other polarimetric detectors, which has not been evaluated yet.

The statistical properties of the DoP shall be obtained in order to assess the detection performance. Based on bivariate complex Gaussian distributions with zero means, the PDFs of the estimated DoP are derived from the joint distributions of the Stokes vector in [13]. They can also be derived from the joint eigenvalue distributions of the polarization covariance matrix in [14], where the polarization covariance matrix is a complex central Wishart matrix. However, when considering a real target embedded in clutter, the complex Gaussian with zero means will be a rough approximation. A non-zero-mean complex Gaussian radar model has been established in [5]. Thus, the DoP statistic, which is used for describing echoes including both targets and clutter, can be obtained from this model.

In the following section, a basic polarimetric radar measurement model is introduced, and the echoes for both target and clutter are described by statistical models. With an assumption of a distribution with non-zero means, the polarization covariance matrix of the range cell including targets follows a complex noncentral Wishart distribution. Based on the PDF of joint eigenvalues of the PCM, the statistics of the DoP for the range cells with or without targets are derived in Section 3. Using the DoP as a detection statistic, a new detector is designed in Section 4. The theoretical performance of the proposed detector is analyzed in this section. Then, in Section 5 the application of the DoP detector and some other typical polarimetric detectors against real data are presented. The performance improvement of the new proposed detector can be demonstrated by the comparison results. Concluding components are provided in Section 6.

2. RADAR RETURN MODEL

We consider that a target with deterministic PSM is being illuminated by polarimetric radar, which has the ability of dual-polarization simultaneous reception (i.e., horizontal and vertical reception). The radar returns also include clutter signals surrounding the target echoes. Specially, the clutter in the main beam of the radar is mainly considered in this paper. Thereupon, the radar return corresponding to the range cell that the target exists can be established as the following model:

$$H_1 : \mathbf{x} = \mathbf{S}\mathbf{h}_t a + \mathbf{c} + \mathbf{n} \quad (1)$$

where H_1 denotes the target-present hypothesis,

$$\mathbf{S} = \begin{bmatrix} s_{hh} & s_{hv} \\ s_{vh} & s_{vv} \end{bmatrix}$$

is the 2×2 PSM of the target, which represents the polarization change of the transmitted signal; \mathbf{h}_t is the 2×1 polarization Jones vector of the transmitted electromagnetic wave; a includes the transmitted radar waveform and information of the range and Doppler. The second term in the right side of Equation (1) represents the clutter signals. \mathbf{n} is the noise in each polarimetric channel. We assume that the target exists in only one range cell. Then in other range cells, the signal model satisfies target-free hypothesis H_0 , which is given by

$$H_0 : \mathbf{x} = \mathbf{c} + \mathbf{n}. \quad (2)$$

In the clutter-only range cell, the radar returns consist of the reflections from large number of random scattering points. Then, the measurement data \mathbf{x} in H_0 case follows the bivariate complex Gaussian distribution with zero means

$$f_{H_0}(\mathbf{x}) = \frac{1}{\pi^2 |\boldsymbol{\Sigma}|} \exp(-\mathbf{x}^H \boldsymbol{\Sigma}^{-1} \mathbf{x}), \quad \mathbf{x} \in \mathbb{C}^2 \quad (3)$$

where H denotes the Hermitian transpose, $\boldsymbol{\Sigma}$ the covariance matrix, and $|\boldsymbol{\Sigma}|$ the determinant of $\boldsymbol{\Sigma}$.

On the other hand, when both the target and clutter exist in the range cell under test, \mathbf{x} in H_1 case is a non-zero-mean complex Gaussian random vector. In this case, the PSM, the relative position and attitude of the target can be considered as constant during pulses in the radar dwell. Then the mean vector of \mathbf{x} is defined as

$$\mathbf{s} \triangleq E(\mathbf{x}) = \mathbf{S}\mathbf{h}_s a$$

and covariance matrix is

$$\mathbf{\Sigma} = E \left[(\mathbf{x} - \mathbf{s})(\mathbf{x} - \mathbf{s})^H \right]$$

then its PDF can be obtained from [15] as

$$f_{H_1}(\mathbf{x}) = \frac{1}{\pi^2 |\mathbf{\Sigma}|} \exp \left\{ -(\mathbf{x} - \mathbf{s})^H \mathbf{\Sigma}^{-1} (\mathbf{x} - \mathbf{s}) \right\}, \quad \mathbf{x} \in \mathbb{C}^2 \quad (4)$$

Assuming that the radar dwell duration consists of K pulses, we can obtain K observation samples for radar returns in each range cell. The k th observation sample for a range cell is defined as a Jones vector $\mathbf{x}_k = [x_{H,k} \ x_{V,k}]^T$. H and V in the subscript denote a set of horizontal and vertical reception polarization basis. Obviously, the random vector \mathbf{x}_k follows the PDF in (4) or (3) as the target exists or not. It is worthwhile to notice that $f_{H_1}(\mathbf{x}) \rightarrow f_{H_0}(\mathbf{x})$ as $\mathbf{s} \rightarrow 0$, when comparing formula (4) and (3).

3. STATISTICS OF THE DOP

It is well known that the target detection problem is to make decision of choosing hypothesis between H_0 and H_1 . In this section, we introduce the degree of polarization as a novel detection statistic. In order to design the target detector by using this statistic, the statistical properties of the DoP should be analyzed in different hypothesis test situations.

As we know, the degree of polarization can be used to characterize the polarization state of the partially polarized waves. We can obtain this parameter from the Stokes vector or polarization covariance matrix. The latter is considered in this paper. Then, the DoP p can be defined in [8] as

$$p \triangleq \frac{\sqrt{\text{tr}(\mathbf{\Sigma})^2 - 4|\mathbf{\Sigma}|}}{\text{tr}(\mathbf{\Sigma})} = \frac{\eta_2 - \eta_1}{\eta_2 + \eta_1} \quad (5)$$

where $\text{tr}(\mathbf{\Sigma})$ denotes the trace of $\mathbf{\Sigma}$, η_1 and η_2 ($\eta_2 \geq \eta_1$) are the eigenvalues of $\mathbf{\Sigma}$. Since we have no prior knowledge about the covariance matrix in real application, it should be estimated from the measurement data. According to the definition of the PCM in [7], the estimation of the PCM $\hat{\mathbf{\Sigma}}$ can be generated from a set of observation samples $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_K$ as follows:

$$\hat{\mathbf{\Sigma}} \triangleq \frac{1}{K} \sum_{k=0}^{K-1} \mathbf{x}_k \mathbf{x}_k^H \quad (6)$$

where K is the integrated number of samples. If the eigenvalues of $\hat{\mathbf{\Sigma}}$ are $\hat{\eta}_1$ and $\hat{\eta}_2$, then the sample estimate of the DoP is defined as

$$\hat{p} \triangleq \frac{\hat{\eta}_2 - \hat{\eta}_1}{\hat{\eta}_2 + \hat{\eta}_1} \quad (7)$$

As we discussed at the end of last section, the hypothesis H_0 can be considered as a special case of H_1 . Since \mathbf{x}_k follows the PDF in (4), let $\mathbf{\Xi} = \sum_{k=0}^{K-1} \mathbf{x}_k \mathbf{x}_k^H = K \hat{\mathbf{\Sigma}}$, then the matrix $\mathbf{\Xi}$ follows the complex noncentral Wishart distribution with K degrees of freedom, when $K \geq 2$. The PDF of $\mathbf{\Xi}$ is given in [16] as

$$f_W(\mathbf{\Xi}) = e^{-\text{tr}(\mathbf{\Theta})} {}_0\tilde{F}_1(K; \mathbf{\Theta} \mathbf{\Sigma}^{-1} \mathbf{\Xi}) f_W^C(\mathbf{\Xi}) \quad (8)$$

where $\mathbf{\Theta} \triangleq \mathbf{\Sigma}^{-1} \mathbf{s} \mathbf{s}^H$ is the non-centrality matrix, and $\text{tr}(\mathbf{\Theta})$ designates the trace of $\mathbf{\Theta}$. ${}_0\tilde{F}_1(\cdot; \cdot)$ in Eq. (8) is the Bessel-type hypergeometric function, a special case of the generic hypergeometric function of matrix argument in the complex field, defined as a series of complex zonal polynomials in [16]. $f_W^C(\mathbf{\Xi})$ is the complex central Wishart distribution, whose PDF takes the form

$$f_W^C(\mathbf{\Xi}) = \frac{|\mathbf{\Xi}|^{K-2}}{\pi \Gamma(K) \Gamma(K-1) |\mathbf{\Sigma}|^K} e^{[-\text{tr}(\mathbf{\Sigma}^{-1} \mathbf{\Xi})]} \quad (9)$$

where $\Gamma(\cdot)$ is the standard Gamma function.

3.1. Joint Eigenvalue Distributions of Polarization Covariance Matrices

Suppose that $\mathbf{\Lambda} = \text{diag}(\lambda_1, \lambda_2)$ and $\mathbf{M} = \text{diag}(\mu_1, \mu_2)$ are the eigenvalue matrices of $\mathbf{\Sigma}^{-1}\mathbf{\Xi}$ and $\mathbf{\Theta}$, respectively. The entries of the two diagonal matrices $\lambda_2 \geq \lambda_1$ and $\mu_2 \geq \mu_1$ are real, since $\mathbf{\Sigma}^{-1}\mathbf{\Xi}$ and $\mathbf{\Theta}$ are both 2×2 Hermitians. From James' results in [16], the joint distribution of the ordered eigenvalues (λ_1, λ_2) corresponding to $\mathbf{\Sigma}^{-1}\mathbf{\Xi}$ can be written as

$$f(\lambda_1, \lambda_2) = \frac{e^{(-\text{tr}M)}e^{(-\text{tr}\Lambda)}}{\Gamma(K)\Gamma(K-1)} {}_0\tilde{F}_1(K; \mathbf{M}, \mathbf{\Lambda}) (\lambda_1\lambda_2)^{K-2} (\lambda_2 - \lambda_1)^2 \quad (10)$$

where ${}_0\tilde{F}_1(\cdot; \cdot, \cdot)$ is the hypergeometric function of 2 matrix arguments. Utilizing the theorem 4.2 in [17], the hyper-geometric function in Eq. (10) can be extended as

$${}_0\tilde{F}_1(K; \mathbf{M}, \mathbf{\Lambda}) = \frac{\Gamma(K) |({}_0F_1(K-1; \lambda_i\mu_j))_2|}{\Gamma(K-1)(\mu_2 - \mu_1)(\lambda_2 - \lambda_1)} \quad (11)$$

where $|(f(i, j))_2|$ denotes the determinant of a 2×2 matrix, $f(i, j)$ the matrix element in the i th row and j th column, and ${}_0F_1(\cdot; \cdot)$ a special case of the generalized hypergeometric function defined as [18]

$${}_pF_q(a_1, a_2, \dots, a_p; b_1, b_2, \dots, b_q; z) \triangleq \sum_{\tau=0}^{\infty} \frac{(a_1)_\tau (a_2)_\tau \dots (a_p)_\tau z^\tau}{(b_1)_\tau (b_2)_\tau \dots (b_q)_\tau \tau!} \quad (12)$$

where $(\alpha)_n = \Gamma(\alpha + n)/\Gamma(\alpha)$. In terms of the denominator in (11), $\lambda_1 \neq \lambda_2$ and $\mu_1 \neq \mu_2$ are required to ensure the formula meaningfully. However, because of the presence of the term $(\lambda_2 - \lambda_1)^2$ in Eq. (10), the requirement that λ_1 should be different with λ_2 can be relaxed. Substituting Eq. (11) into Eq. (10), the joint distribution of (λ_1, λ_2) is given by

$$f_{fr}(\lambda_1, \lambda_2) = \frac{e^{-(\mu_1 + \mu_2 + \lambda_1 + \lambda_2)} (\lambda_1\lambda_2)^{K-2} (\lambda_2 - \lambda_1)}{\Gamma(K-1)^2 (\mu_2 - \mu_1)} \cdot |({}_0F_1(K-1; \lambda_i\mu_j))_2|. \quad (13)$$

In terms of the definition of the non-centrality matrix $\mathbf{\Theta}$, it is the product of the inverse of the covariance matrix $\mathbf{\Sigma}$ and the outer product \mathbf{ss}^H of the mean vector \mathbf{s} . Due to the rank-one term \mathbf{ss}^H , it is worthwhile to notice that $\mathbf{\Theta}$ has at most rank one, i.e., it has at least one zero eigenvalue. First, we consider the case that only one eigenvalue of $\mathbf{\Theta}$ is zero, i.e., $\mu_2 > \mu_1 = 0$, we can call it rank-1 case. According to $\lim_{z \rightarrow 0} {}_0F_1(a; z) = 1$ and taking the limit of (13) as $\mu_1 \rightarrow 0$, the joint PDF of (λ_1, λ_2) can be reduced to

$$f_{r1}(\lambda_1, \lambda_2) = \frac{e^{-(\lambda_1 + \lambda_2 + \mu_2)} (\lambda_1\lambda_2)^{K-2} (\lambda_2 - \lambda_1)}{\Gamma(K-1)^2 \mu_2} \cdot [{}_0F_1(K-1; \lambda_2\mu_2) - {}_0F_1(K-1; \lambda_1\mu_2)]. \quad (14)$$

Second, two eigenvalues of $\mathbf{\Theta}$ are both zero in the so-called rank-0 case, i.e., $\mu_2 = \mu_1 = 0$. Applying L'Hospital's rule in taking the limit of the joint PDF in Eq. (14) and using

$$\frac{\partial {}_0F_1(n; az)}{\partial z} = \frac{a {}_0F_1(n+1; az)}{n} \quad (15)$$

yields

$$f_{r0}(\lambda_1, \lambda_2) = \lim_{\mu_2 \rightarrow 0} f_{r1}(\lambda_1, \lambda_2; \mu_2) = \frac{e^{-(\lambda_1 + \lambda_2)} (\lambda_1\lambda_2)^{K-2} (\lambda_2 - \lambda_1)^2}{\Gamma(K)\Gamma(K-1)}. \quad (16)$$

The joint eigenvalue PDFs of $\mathbf{\Sigma}^{-1}\mathbf{\Xi}$ have been expressed in closed forms for different cases. It is easy to found that rank-1 and rank-0 cases are just corresponding to H_1 and H_0 hypotheses, respectively. PDFs of the estimated DoP in each case can be derived through variable transformation and marginal PDF calculation next.

3.2. PDFs of the Estimated Degree of Polarization

Some important special cases will be focused in the sequel, in which the polarization components of clutter in horizontal and vertical polarimetric basis are supposed mutual independent. Then the

covariance matrix can be denoted as $\mathbf{\Sigma} = \sigma^2 \mathbf{I}_2$, where \mathbf{I}_2 is the 2×2 identity matrix and σ^2 the clutter power in each polarization basis.

In terms of the definition of $\mathbf{\Xi}$ and $\mathbf{\Sigma}$, the maximum likelihood estimator of the true polarization covariance matrix can be denoted as $\hat{\mathbf{\Sigma}} = \sigma^2 (\mathbf{\Sigma}^{-1} \mathbf{\Xi}) / K$. Then the relations between eigenvalues of $\hat{\mathbf{\Sigma}}$ and $\mathbf{\Sigma}^{-1} \mathbf{\Xi}$ can be established as $\hat{\eta}_1 = \sigma^2 \lambda_1 / K$ and $\hat{\eta}_2 = \sigma^2 \lambda_2 / K$. Substituting the relations into Eq. (7), the estimated DoP is given by

$$\hat{p} = \frac{\frac{\sigma^2}{K} \lambda_2 - \frac{\sigma^2}{K} \lambda_1}{\frac{\sigma^2}{K} \lambda_2 + \frac{\sigma^2}{K} \lambda_1} = \frac{\lambda_2 - \lambda_1}{\lambda_2 + \lambda_1}. \quad (17)$$

We consider the rank-1 case at first. Given the joint PDF of (λ_1, λ_2) in Eq. (14), we can transform to the new variables $\hat{p} = (\lambda_2 - \lambda_1) / (\lambda_2 + \lambda_1)$ and $q = \lambda_2 - \lambda_1$ by the standard formula for the change of variables in PDFs. For this transformation, the Jacobian can be denoted as

$$J \left(\frac{\lambda_1, \lambda_2}{\hat{p}, q} \right) = q / (2\hat{p}^2). \quad (18)$$

Then we can find the joint distribution of (\hat{p}, q)

$$f_{r1}(\hat{p}, q) = \frac{e^{-\mu_2} (1 - \hat{p}^2)^{K-2}}{2^{2K-3} \Gamma(K-1)^2 \mu_2 \hat{p}^{2K-2}} e^{-\frac{q}{\hat{p}} q^{2K-2}} \cdot \left[{}_0F_1 \left(K-1; \frac{(1+\hat{p})\mu_2}{2\hat{p}} q \right) - {}_0F_1 \left(K-1; \frac{(1-\hat{p})\mu_2}{2\hat{p}} q \right) \right]. \quad (19)$$

Integration over q from 0 to ∞ using the Equation (7.522-9) in [18] yields

$$f_{r1}(\hat{p}) = \frac{\Gamma(2K-1) e^{-\mu_2} (1 - \hat{p}^2)^{K-2} \hat{p}}{2^{2K-3} \Gamma(K-1)^2 \mu_2} \cdot \left[{}_1F_1 \left(2K-1; K-1; \frac{1+\hat{p}}{2} \mu_2 \right) - {}_1F_1 \left(2K-1; K-1; \frac{1-\hat{p}}{2} \mu_2 \right) \right] \quad (20)$$

where ${}_1F_1(\cdot; \cdot; \cdot)$ is the confluent hypergeometric function, also a special case of the generalized hypergeometric series defined in Eq. (12).

We define SCR as ratio of the target echoes power to the total power of clutter. According to the definition of $\mathbf{\Theta}$ and $\mu_1 = 0$, we have

$$\mu_2 = \text{tr}(\mathbf{\Theta}) = \frac{\|\mathbf{s}\|^2}{\sigma^2} = 2\text{SCR}. \quad (21)$$

where $\|\mathbf{s}\|^2$ is the power of the target echoes, and $2\sigma^2$ the total clutter power. From Eqs. (20) and (21), we can easily find that the PDF of \hat{p} in rank-1 case has relation with not only the number of samples but SCR as well.

In rank-0 case, similar variable transformation from (λ_1, λ_2) to (\hat{p}, q) and marginal integration over q can be done to Eq. (10). Then the PDF $f_{r0}(\hat{p})$ can be given as

$$f_{r0}(\hat{p}) = \frac{\Gamma(2K)(1 - \hat{p}^2)^{K-2} \hat{p}^2}{2^{2K-3} \Gamma(K) \Gamma(K-1)}. \quad (22)$$

The result in Eq. (22) is identical to the PDF in [10], as the real DoP $P \rightarrow 0$. Since the eigenvalues of $\mathbf{\Theta}$ are both zero in rank-0 case, none of the deterministic polarization components exist in radar reception. Consequently, the polarization covariance matrix degrades to the central Wishart distribution, and the real DoP approximates to 0. As can be noted in Eq. (22), the PDF of \hat{p} in this case is merely the function of the number of samples.

4. DETECTION TEST AND PERFORMANCE

The novel detection statistic is defined as

$$T_{DoP} = \hat{p}. \quad (23)$$

An illustration of the PDFs for the DoP hypothesis test is shown in Fig. 1. Then the detector decides H_1 , if $T_{DoP} > \gamma$. Otherwise, H_0 is chosen. γ is a threshold of DoP, which is plotted by the dotted

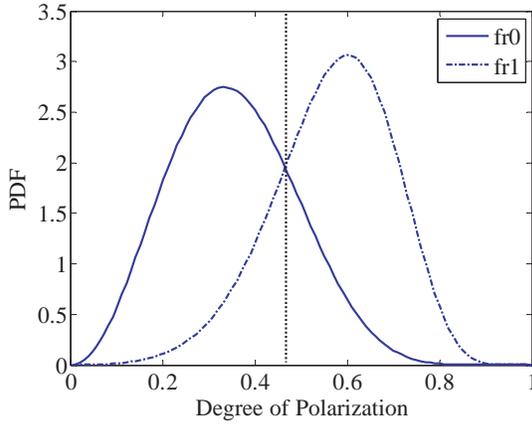


Figure 1. PDFs for the DoP hypothesis test.

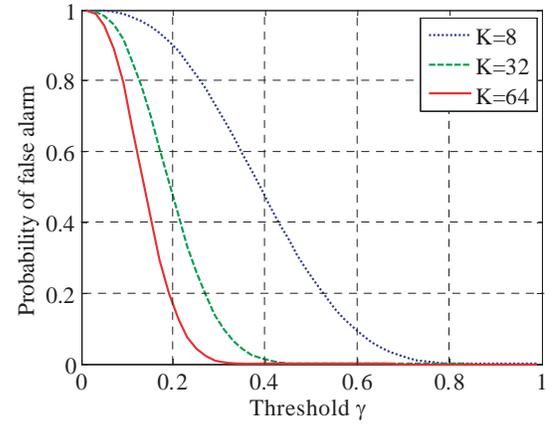


Figure 2. Probability of false alarm of the detector as a function of the threshold, for different sample number K .

lines in Fig. 1. According to the results obtained in last section, the detection statistic is distributed as follows:

$$T_{DoP} \sim \begin{cases} f_{r0}(\hat{p}) & \text{under } H_0 \\ f_{r1}(\hat{p}) & \text{under } H_1 \end{cases} \quad (24)$$

Thus, the probability of false alarm can be expressed as follows:

$$P_{FA} = P(T_{DoP} > \gamma | H_0) = \int_{\gamma}^1 f_{r0}(\hat{p}) d\hat{p} = 1 - \int_0^{\gamma} f_{r0}(\hat{p}) d\hat{p} = 1 - \frac{B_{\gamma^2}(K-1, \frac{3}{2})}{B(K-1, \frac{3}{2})} \quad (25)$$

where $B(a, b)$ and $B_{\alpha}(a, b)$ are beta function and incomplete beta function, respectively. It is worthwhile to mention that the expression for P_{FA} has only relationship with the threshold γ and sample number K . Consequently, the DoP detector is a CFAR test. Fig. 2 shows the probability of false alarm as a function of the threshold γ for sample number $K = 8, 32$ and 64 . As can be easily noted that P_{FA} decreases as γ increases. In addition, the P_{FA} curve becomes steeper as the sample number K increases.

According to the required P_{FA} , the threshold γ can be determined by Eq. (25). Then the detection probability can be obtained by

$$P_D = \int_{\gamma}^1 f_{r1}(\hat{p}) d\hat{p} = Q_{f_{r1}}(\gamma) \quad (26)$$

where Q denotes the right-tail probability function.

To evaluate the performance improvement, the DoP detector is compared with some other typical polarimetric detectors. The single-polarization generalized likelihood ratio (SP-GLR) detector was derived by using only one polarimetric channel (HH for example) [6]. Polarization-space-time generalized likelihood ratio (PST-GLR) detector was established in [3]. These two detectors both need secondary data to estimate the clutter covariance matrix. Multi-polarization generalized likelihood ratio (MP-GLR) detector was formulated assuming that the clutter is inhomogeneous [5]. The secondary data are unnecessary for this detector. However, it can only detect the static or slow target.

For comparison, we set the desired $P_{FA} = 10^{-2}, 10^{-3}$ and 10^{-4} , respectively. For different values of sample number K , the threshold can be easily found based on Fig. 2. Then, we obtain P_D by calculating the right-tail probability function in Eq. (26) as a function of SCR. The probabilities of detection P_D of all the detectors mentioned above as a function of SCR are plotted in Figs. 3(a), (b) and (c). Specifically, the newly proposed DoP detector is shown as dash-dotted lines. The SP-GLR, MP-GLR and PST-GLR detectors are plotted as solid lines, dashed lines and plus-mark lines, respectively. The three subfigures in Fig. 3 correspond to different values of P_{FA} . The figures show that the detection performance of the DoP detector and MP-GLR is significantly better than that of the other two detectors for this case.

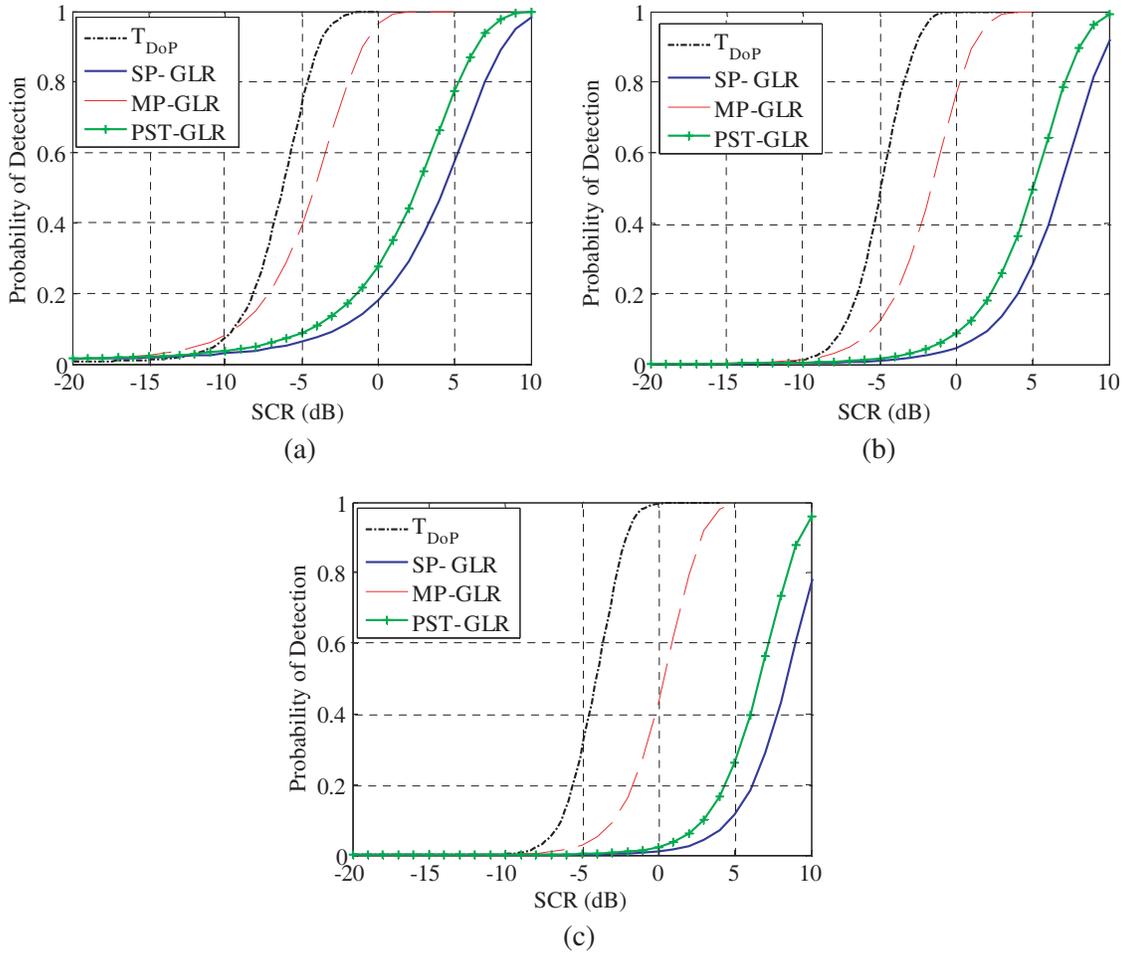


Figure 3. Detection probabilities of the polarimetric detectors as a function of the SCR for different values of P_{FA} . (a) $P_{FA} = 10^{-2}$. (b) $P_{FA} = 10^{-3}$. (c) $P_{FA} = 10^{-4}$.

Furthermore, since no reference data from different range cells are needed for the DoP and MP-GLR detectors, they are both suitable for the nonhomogeneous clutter environment.

In terms of the DoP detector, the detection performance is independent of the secondary sample number, which can be noticed from the three subfigures in Fig. 3. We can also compare the detection performances of the DoP detector with that of the MP-GLR as a function of SCR. For this case, we assume that the target is not static. Fig. 3 shows that for this moving-target case, the detection probability of the DoP detector exhibits a significant improvement compared to that of the MP-GLR detector, when the total SCRs in the two detectors are the same. In addition, more obvious improvement can be observed as the SCR increases. Therefore, target detection using the DoP statistics seems more convenient in real polarimetric data. To support this claim, we will evaluate the detection performance by applying all the four detectors to real data.

5. APPLICATION TO REAL DATA

The data collected by the IPIX radar of McMaster University were used to assess the detector performance in real environment [19]. IPIX radar has the ability of dual-polarized and frequency agile operation. The data we used were recorded on November 18, 1993, at Osborne Head Gunnery Range (OHGR). The specific data parameters are listed in Table 1. The measurement range of the data is 2649 m–2844 m, in which a target exists. Its location is about 2655 m from the radar set and azimuth

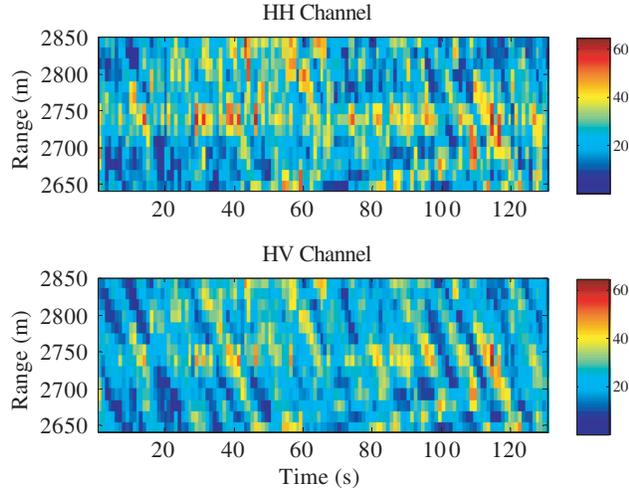


Figure 4. Time-range images of the IPIX radar dataset 1 in HH and HV channels.

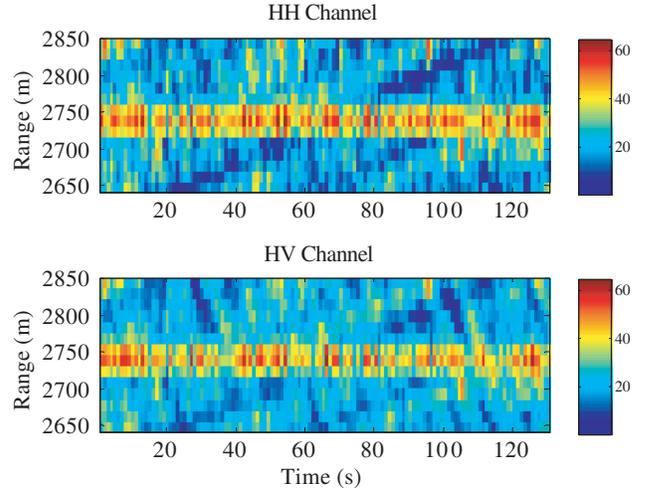


Figure 5. Time-range images of the IPIX radar dataset 2 in HH and HV channels.

Table 1. Parameters of IPIX radar datasets.

Name	Dataset 1	Dataset 2
Recording Date	1993.11.18	1993.11.18
Time	02:36	16:26
Target Location	2655 m, 170 deg	2655 m, 170 deg
Range Resolution	30 m	30 m
Wave Height	1.4 m	0.9 m
Wind Speed	26 km/h	17 km/h

angle 170 degrees. Figs. 4 and 5 show the magnitude of the echoes from both the target and the sea clutter in polarimetric channels HH and VH .

Figure 4 corresponds to the dataset in the first row of Table 1. We can observe that the clutter reflects similar energy to the target in both HH and VH polarimetric channels. The dataset in the second row of the table is shown in Fig. 5, where the reflection of clutter is much lower than the target because of the slower wind speed and smaller sea wave.

Although the exact ratio of the target to clutter is unknown previously, to compare the detection performance, we ran the DoP, SP-GLR, PST-GLR and MP-GLR detectors against the former datasets. Then we plotted the P_D versus P_{FA} curves in Fig. 6, which are also called receiver operating characteristic (ROC). Since the target location in the datasets was a priori knowledge, we divided the detection map into two regions: target and clutter. Then we can calculate the detection probability of the data in the target region by counting the number of pixels in which the statistic is larger than the given threshold as P_D . The false alarm probability P_{FA} can be computed from the data in the clutter region by the same way. A pair (P_{FA}, P_D) corresponds to a point on the ROC curve.

Figure 6 shows the ROC curves for the four detectors: (a) for dataset 1 and (b) for dataset 2. The different pairs of (P_{FA}, P_D) are derived by changing the threshold values. The dash-dotted lines still represent performance of the DoP detector. The SP-GLR, MP-GLR and PST-GLR detectors are plotted as solid lines, dashed lines and plus-mark lines, respectively. In both Figs. 6(a) and (b), we can notice that our detector T_{DoP} usually outperforms the other three detectors. In addition, compared with the SP-GLR detector, the DoP detection performance improvement in Fig. 6(b) is more obvious than that in Fig. 6(a) because of the higher SCR. This is consistent with the conclusion we drew in last section.

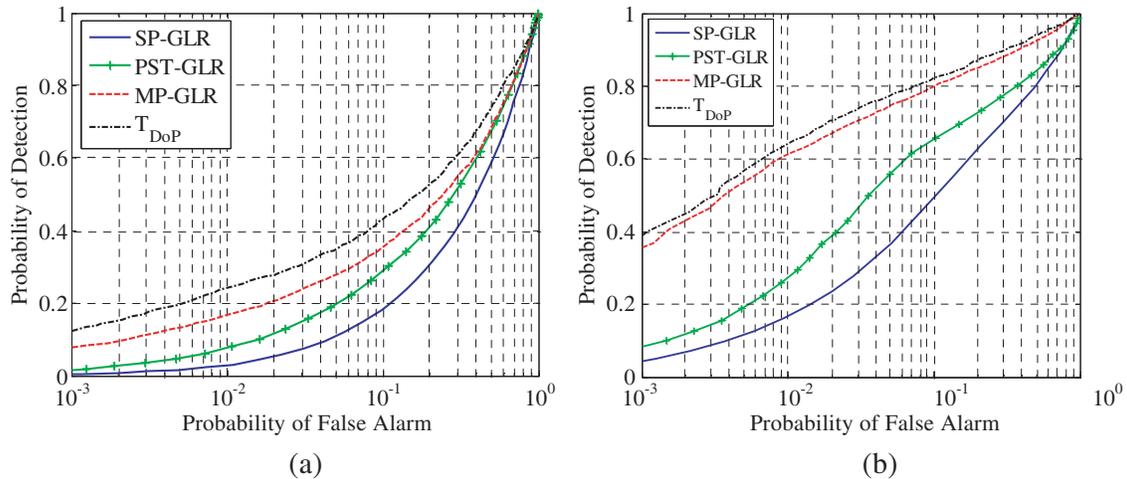


Figure 6. Detector performance for real data collected by IPIX radar on November 18, 1993. (a) Dataset 1. (b) Dataset 2.

6. CONCLUSION

In this paper, we have designed a novel target detector by using statistics of the estimated DoP. We generated the polarimetric radar model at first which includes both target and clutter echoes. Based on the radar model, the PDFs of the estimated DoP for the radar reception were derived from joint eigenvalue distributions of the complex noncentral Wishart matrix. To evaluate the detection performance, we analyzed the probabilities of false alarm and detection. The performance of the new detector was also compared with the SP-GLR, PST-GLR and MP-GLR detectors. At last, we applied the detectors in real collected data. For the proposed detector, the compared results demonstrated an obvious performance improvement with respect to the other typical polarimetric detectors.

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